

# HKN ECE 313 EXAM 2 REVIEW SESSION

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## EXAM 2 TOPICS

- Continuous-type Random Variables (mean and variance of CRVs)
- Uniform Distribution
- Exponential Distribution
- Poisson Process
- Linear Scaling of PDFs
- Gaussian Distribution
- ML Parameter Estimation for Continuous Random Variables
- Functions of a random variable
- Failure Rate Functions
- Binary Hypothesis Testing
- Joint CDFs, PMFs, and PDFs
- Independence of Random Variables
- Distributions of sums of random variables

## CONTINUOUS-TYPE RANDOM VARIABLES

- Cumulative Distribution Functions (CDFs)
  - Must be nondecreasing
  - Must go to 0 at  $-\infty$  and 1 at  $+\infty$
  - Must be right continuous
- $F_X(c) = \int_{-\infty}^c f_X(u)du$
- $P\{X = a\} = 0$
- $P\{a < X \leq b\} = F_X(b) - F_X(a) = \int_a^b f_X(u)du$
- $E[X] = \mu_x = \int_{-\infty}^{+\infty} u f_X(u)du$

# UNIFORM DISTRIBUTION

- Uniform(a,b):
- pdf:  $f(u) = \begin{cases} \frac{1}{b-a} & a \leq u \leq b \\ 0 & \text{else} \end{cases}$
- mean:  $\frac{a+b}{2}$
- variance:  $\frac{(b-a)^2}{12}$

# EXPONENTIAL DISTRIBUTION

- Exponential( $\lambda$ ): limit of scaled geometric random variables
- pdf:  $f(t) = \lambda e^{-\lambda t} \quad t \geq 0$
- mean:  $\frac{1}{\lambda}$
- variance:  $\frac{1}{\lambda^2}$
- Memoryless Property
  - $P\{T \geq s + t \mid T \geq s\} = P\{T \geq t\} \quad s, t \geq 0$

# POISSON PROCESS

- Poisson process is used to model the number of counts in a time interval. Similar to how the exponential distribution is a limit of the geometric distribution, the Poisson Process is the limited of the Bernoulli process.
- *If we have a rate  $\lambda$  and time interval  $t$ , the number of counts  $\sim \text{Poisson}(\lambda t)$*
- Disjoint intervals, e.g. 0 to 2s and 2 to 3s, are independent

## LINEAR SCALING OF PDFS

- If  $Y = aX + b$ :
- $E[Y] = aE[X] + b$
- $Var(Y) = a^2Var(X)$
- $f_y(v) = f_x\left(\frac{v-b}{a}\right) \frac{1}{a}$

# GAUSSIAN DISTRIBUTION

- Gaussian (or Normal) Distribution  $\sim N(\mu, \sigma^2)$ 
  - Standard Gaussian,  $\hat{X}: \mu = 0, \sigma = 1$
- pdf:  $f(u) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(u-\mu)^2}{2\sigma^2}}$
- mean:  $\mu$
- variance:  $\sigma^2$
- If we standardize our Gaussian, where:  $\hat{X} = \frac{X-\mu}{\sigma}$ 
  - $\Phi(c) = \int_{-\infty}^c f(u)du$
  - $Q(c) = 1 - \Phi(c) = \int_c^{\infty} f(u)du$



# ML PARAMETER ESTIMATION

- Suppose we have a random variable with a given distribution/pdf that depends on a parameter,  $\theta$ . By taking trials of the random variable, we can estimate  $\theta$  by finding the value that maximizes the likelihood of the observed event,  $\hat{\theta}_{ML}$ .
- There are a few ways we can find  $\hat{\theta}_{ML}$ 
  - Take derivative of provided pmf and set it equal to zero (maximization)
  - Observe the intervals where the likelihood increases and decreases, and find the maximum between these intervals
  - Intuition!

# FUNCTIONS OF RANDOM VARIABLES

- Suppose  $Y = g(X)$ , and we want to be able to describe the distribution of  $Y$
- Step 1: Identify the support of  $X$ . Sketch the pdf of  $X$  and  $g$ . Identify the support of  $Y$ . Determine whether  $Y$  is a Continuous or Discrete RV
- Step 2 (for CRV): Use the definition of the CDF to find the CDF of  $Y$ :
  - $F_Y(c) = P\{Y \leq c\} = P\{g(X) \leq c\}$
- Step 2 (for DRV): Find the pmf of  $Y$  directly using the definition of the pmf
  - $p_Y(v) = P\{Y = v\} = P\{g(X) = v\}$
- Step 3 (for CRV): Differentiate the CDF of  $Y$  in order to find the pdf of  $Y$

## GENERATING A RV WITH A SPECIFIED DISTRIBUTION

- We can generate any distribution by applying a function to a uniform distribution
- This function should be the inverse of the CDF of the desired distribution
- Ex: if we want an exponential distribution,
  - $F_X(c) = 1 - e^{-c} = u$ ; then find  $F_X^{-1}(c)$

# FAILURE RATE FUNCTIONS

- We can assess the probability of a failure in a system through a failure rate function,  $h(t)$ .
- $F(t) = 1 - e^{-\int_0^t h(s)ds}$
- Two popular failure rate functions:
  - Consistent lifetime
  - “Bath tub”

# BINARY HYPOTHESIS TESTING

- Similar to BHT with Discrete RVs
- Maximum Likelihood (ML) Rule
  - $\Lambda(k) = \frac{f_1(k)}{f_0(k)}$
  - $\Lambda(k) = \begin{cases} > 1 \text{ declare } H_1 \text{ is true} \\ < 1 \text{ declare } H_0 \text{ is true} \end{cases}$
- Maximum a Posteriori (MAP) Rule
  - Prior probabilities:  $\pi_0 = P(H_0), \pi_1 = P(H_1)$
  - $H_1$  true if  $\pi_1 f_1(k) > \pi_0 f_0(k)$ , same as  $\Lambda(k) = \frac{f_1(k)}{f_0(k)} > \tau$  where  $\tau = \frac{\pi_0}{\pi_1}$
- Probabilities of False Alarm and Miss
  - $p_{\text{false alarm}} = P(\text{Say } H_1 \mid H_0 \text{ is true})$
  - $p_{\text{miss}} = P(\text{Say } H_0 \mid H_1 \text{ is true})$
  - $p_e = \pi_1 p_{\text{miss}} + \pi_0 p_{\text{false alarm}}$

# JOINT CDF, PMF, AND PDF

## DISCRETE RANDOM VARIABLES

- Joint CDF
  - $F_{X,Y}(u_o, v_o) = P\{X \leq u_o, Y \leq v_o\}$
- Joint PMF
  - $p_{X,Y}(u_o, v_o) = P\{X = u_o, Y = v_o\}$
- Marginal PMFs
  - $p_X(u) = \sum_j p_{X,Y}(u, v_j)$
  - $p_Y(v) = \sum_i p_{X,Y}(u_i, v)$
- Conditional PMFs
  - $p_{(Y|X)}(v|u_o) = P(Y = v|X = u_o) = \frac{p_{X,Y}(u_o, v)}{p_X(u_o)}$

## CONTINUOUS RANDOM VARIABLES

- Joint CDF
  - $F_{X,Y}(u_o, v_o) = P\{X \leq u_o, Y \leq v_o\}$
- Joint PDF
  - $F_{X,Y}(u_o, v_o) = \iint f_{X,Y}(u, v) dv du$
- Marginal PDFs
  - $f_X(u) = \int_{-\infty}^{+\infty} f_{X,Y}(u, v) dv$
  - $f_Y(v) = \int_{-\infty}^{+\infty} f_{X,Y}(u, v) du$
- Conditional PDFs
  - $f_{(Y|X)}(v|u_o) = \frac{f_{X,Y}(u_o, v)}{f_X(u_o)}$

# INDEPENDENCE OF JOINT DISTRIBUTIONS

- We can check independence in a joint distribution in a couple ways:
- $f_{X,Y}(u, v) = f_X(u)f_Y(v)$
- The support of  $f_{X,Y}(u, v)$  is a *product set*
  - Product set must have the swap property, which is satisfied if:
    - $(a, b)$  and  $(c, d) \in \text{support}(f_{X,Y})$ , and then  $(a, d)$  and  $(b, c)$  are also  $\in \text{support}(f_{X,Y})$
  - Checking for a product set is only sufficient to prove *dependence*. Saying that the joint pdf is a product set is not sufficient to check independence

## DISTRIBUTION OF SUMS OF RANDOM VARIABLES

- Suppose we want to find the distribution from the sum of two independent random variables where  $Z = X + Y$
- The pdf or pmf of  $Z$  is the *convolution* of the two pdfs/pmfs



## FA15 PROBLEM 5

5. [18 points] Suppose we are testing if a transistor in a circuit is working properly or not. The voltage we observe at the output is a normal random variable  $X$ . If the transistor is working,  $X$  is distributed according to  $N(1, 1)$ . If the transistor is not working, then  $X$  is distributed according to  $N(-1, 1)$ . There is an 50% chance that the transistor is working. You can express the answers for this problem in terms of  $\Phi$  and  $Q$ .

(a) (6 points) Find  $P\{X \leq 1 | \text{transistor is working}\}$ .

(b) (6 points) Find  $P\{X \geq 1\}$ .

(c) (6 points) Obtain the unconditional pdf of  $X$ ,  $f_X(u)$  for all  $u$ .

## FA 15 PROBLEM 2

2. [22 points] Let  $N_t$  be a Poisson process with rate  $\lambda > 0$ .

(a) (4 points) Obtain  $P\{N_3 = 5\}$ .

(b) (6 points) Obtain  $P\{N_7 - N_4 = 5\}$  and  $E[N_7 - N_4]$ .

(c) (6 points) Obtain  $P\{N_7 - N_4 = 5 | N_6 - N_4 = 2\}$ .

(d) (6 points) Obtain  $P\{N_6 - N_4 = 2 | N_7 - N_4 = 5\}$ .

## FA14 PROBLEM 3

3. [12 points] Let  $\{N(t), t \geq 0\}$  be a Poisson process with rate  $\lambda$ .

(a) (6 points) Express  $E[N(t)N(t+s)]$ ,  $s, t > 0$  as a function of  $\lambda$ ,  $s$ , and  $t$ .

(b) (6 points) Let  $\lambda = 2$  arrivals/hour and assume that the Poisson process models the arrival of customers into a post office. Find the probability of the following event, which involves three conditions:

(Three customers arrive between 1 and 3pm,  
one customer arrives between 2 and 3pm,  
and one customer arrives between 2 and 4pm.)

## FA15 PROBLEM 6 (MIDTERM I)

6. [20 points] Consider a random variable  $X$  uniformly distributed on the set  $\{1, \dots, n\} \cup \{2n+1, \dots, 3n\}$ , i.e.  $P\{X = k\}$  is constant for  $k = 1, \dots, n, 2n+1, \dots, 3n$ .

- (a) Suppose that  $n$  is unknown but it is observed that  $X = 9$ . Obtain the maximum likelihood estimate of  $n$ .

- (b) Suppose now that it is known that  $n$  can have two different known values,  $n_1$  and  $n_2$ , which gives rise to two hypotheses

$$H_0 : X \in \{1, \dots, n_1\} \cup \{2n_1 + 1, \dots, 3n_1\},$$

$$H_1 : X \in \{1, \dots, n_2\} \cup \{2n_2 + 1, \dots, 3n_2\},$$

where  $n_1 < n_2 < 2n_1$ . Obtain the maximum likelihood decision rule.

- (c) Using the decision rule and distributions from part (b), obtain  $p_{\text{false alarm}}$ .

## SPI5 PROBLEM 4

4. [**16 points**] The two parts of this problem are unrelated.

(a) Let  $X$  be uniformly distributed in  $[0, 1]$ . Find the CDF for

$$Y = 2|X - 1/2|.$$

## FA 15 PROBLEM 4

4. [20 points] Let the joint pdf for the pair  $(X, Y)$  be

$$f_{X,Y}(x, y) = \begin{cases} cxy, & 0 \leq x \leq 1, 0 \leq y \leq 1, x + y \leq 1 \\ 0, & \text{otherwise,} \end{cases}$$

for some constant  $c$ .

(a) (5 points) Compute the marginal  $f_X(x)$ . You can leave it in terms of  $c$ .

(b) (6 points) Obtain the value of the constant  $c$  for  $f_{X,Y}$  to be a valid joint pdf.

(c) (5 points) Obtain  $P\{X + Y < \frac{1}{2}\}$ .

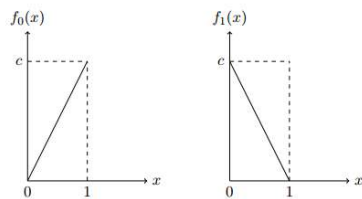
(d) (4 points) Are  $X$  and  $Y$  independent? Explain why or why not.

## FA12 PROBLEM 6

6. [24 points] Suppose random variables  $X$  and  $Y$  have the joint probability density function (pdf):  $f_{XY}(u, v) = \begin{cases} \frac{3}{2}, & u > 0, u^2 < v < 1 \\ 0, & \text{elsewhere} \end{cases}$
- (a) (4 points) Are  $X$  and  $Y$  independent? Explain your answer.
  - (b) (6 points) Determine the marginal pdf of  $X$ ,  $f_X(u)$ .
  - (c) (3 points) For what values of  $u$  is the conditional pdf of  $Y$  given  $X = u$ ,  $f_{Y|X}(v|u)$ , well defined?
  - (d) (4 points) Determine  $f_{Y|X}(v|u)$  for the values of  $u$  for which it is well defined. Be sure to indicate where its value is zero.
  - (e) (7 points) Determine  $P\{Y > X\}$ .

## FA15 PROBLEM 3

3. [20 points] Let  $X$  be a continuous-type random variable taking values in  $[0, 1]$ . Under hypothesis  $H_0$ , the pdf of  $X$  is  $f_0$ ; under hypothesis  $H_1$ , the pdf of  $X$  is  $f_1$ . Both pdfs are plotted below. The priors are known to be  $\pi_0 = 0.6$  and  $\pi_1 = 0.4$ .



- (a) (4 points) Find the value of  $c$ .
- (b) (8 points) Specify the maximum a posteriori (MAP) decision rule for testing  $H_0$  vs.  $H_1$ .
- (c) (8 points) Find the error probabilities  $p_{\text{false alarm}}$ ,  $p_{\text{miss detection}}$  and the average probability of error  $p_e$  for the MAP rule.